

ON THE PROBABLE NUMBER OF JELLY BEANS IN A POUND

BU-252-M

D. S. Robson

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Abstract

If a population of objects is sampled by weight, i.e. if a sample of preassigned weight t is drawn at random then the number N_t of objects in the sample is a random variable. Neglecting excesses over the boundary,

$$P(N_t = n) = F_n(t) - F_{n+1}(t)$$

where $F_n(t)$ is the n -fold convolution of the weight distribution in the population. When t is large relative to the weight of a single object then the Central Limit Theorem gives

$$P(n_1 \leq N_t \leq n_2) \doteq \Phi\left(\frac{t - n_1\mu}{\sigma\sqrt{n_1}}\right) - \Phi\left(\frac{t - (n_2+1)\mu}{\sigma\sqrt{n_2+1}}\right)$$

where μ is the mean weight per object in the population, σ^2 is the variance in weight in the population, and Φ is the standard cumulative normal distribution function.

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In sampling a population of objects the sample size is sometimes determined by total weight rather than total number. For example, in agronomic experiments a predetermined number of grams of seed may be planted in each experimental plot. We consider here the probability distribution of the number of objects N_t in a random sample of predetermined total weight t .

Let $F_n(t)$ denote the n -fold convolution of the probability distribution of the weights of the individual objects in the population,

$$F_n(t) = P(\text{total weight of } n \text{ randomly chosen objects in less than } t).$$

Since the weight of an individual object is non-negative, we then have

$$P(N_t = n) = F_n(t) - F_{n+1}(t).$$

A small difficulty arises in the case $n=0$,

$$\begin{aligned} P(N_t = 0) &= 1 - F_1(t) \\ &= P(\text{wt. of a single randomly chosen object exceeds } t) \end{aligned}$$

and we ignore this boundary problem by assuming that t is very large relative to the weight of a single object, as is the usual case in sampling-by-weight experiments.

This same assumption implies that N_t is stochastically a large number and that we may therefore invoke the Central Limit Theorem to obtain the approximation

$$P(N_t = n) \doteq \Phi \left(\frac{t-n\mu}{\sigma\sqrt{n}} \right) - \Phi \left(\frac{t-(n+1)\mu}{\sigma\sqrt{n+1}} \right)$$

where μ denotes the mean weight per object in the population, σ^2 is the variance in weight among individual objects in the population, and Φ is the standard cumulative normal distribution function. Thus, if z_α is defined by $\Phi(z_\alpha) = 1 - \alpha/2$ then

$$\begin{aligned} P \left\{ \frac{t}{\mu} + \frac{z_\alpha}{2} \left(\frac{\sigma}{\mu} \right)^2 - 1 - \frac{\sigma}{\mu} \sqrt{\frac{tz_\alpha}{\mu} + \left(\frac{\sigma z_\alpha}{2\mu} \right)^2} \leq N_t \right. \\ \left. \leq \frac{t}{\mu} + \frac{z_\alpha}{2} \left(\frac{\sigma}{\mu} \right)^2 + \frac{\sigma}{\mu} \sqrt{\frac{tz_\alpha}{\mu} + \left(\frac{\sigma z_\alpha}{2\mu} \right)^2} \right\} \doteq 1 - \alpha . \end{aligned}$$

Note that another more crude approximation to the distribution of N_t is then given by

$$P(N_t = n) = - \frac{d}{dn} \Phi \left(\frac{t-n\mu}{\sigma\sqrt{n}} \right) .$$